



Slides adapted from Rob Schapire

### Classification: Rademacher Complexity

Machine Learning: Jordan Boyd-Graber University of Colorado Boulder Nothing new ...

- Samples *S* = ((*x*<sub>1</sub>, *y*<sub>1</sub>),...,(*x<sub>m</sub>*, *y<sub>m</sub>*))
- Labels  $y_i = \{-1, +1\}$
- Hypothesis  $h: X \rightarrow \{-1, +1\}$
- Training error:  $\hat{R}(h) = \frac{1}{m} \sum_{i=1}^{m} \mathbb{1}[h(x_i) \neq y_i]$

$$\hat{R}(h) = \frac{1}{m} \sum_{i}^{m} \mathbb{1}\left[h(x_i) \neq y_i\right]$$
(1)

(2)

(3)

(4)

$$\hat{R}(h) = \frac{1}{m} \sum_{i}^{m} \mathbb{1} \left[ h(x_i) \neq y_i \right]$$

$$= \frac{1}{m} \sum_{i}^{m} \begin{cases} 1 & \text{if } (h(x_i, y_i) == (1, -1) \text{ or } (-1, 1) \\ 0 & (h(x_i, y_i) == (1, 1) \text{ or } (-1, -1) \end{cases}$$
(2)

(4)

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$$= \frac{1}{m} \sum_{i}^{m} \frac{1 - y_i h(x_i)}{2}$$
(3)

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$$= \frac{1}{m} \sum_{i}^{m} \frac{1 - y_{i}h(x_{i})}{2}$$

$$= \frac{1}{2} - \frac{1}{2m} \sum_{i}^{m} y_{i}h(x_{i})$$

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(1)
(2)
(3)

Correlation between predictions and labels

$$\hat{R}(h) = \frac{1}{m} \sum_{i}^{m} \mathbb{1} [h(x_{i}) \neq y_{i}]$$

$$= \frac{1}{m} \sum_{i}^{m} \begin{cases} 1 & \text{if } (h(x_{i}, y_{i}) == (1, -1) \text{ or } (-1, 1) \\ 0 & (h(x_{i}, y_{i}) == (1, 1) \text{ or } (-1, -1) \end{cases}$$

$$= \frac{1}{m} \sum_{i}^{m} \frac{1 - y_{i}h(x_{i})}{2}$$

$$= \frac{1}{2} - \frac{1}{2m} \sum_{i}^{m} y_{i}h(x_{i})$$

$$(1)$$

Minimizing training error is thus equivalent to maximizing correlation

$$\arg\max_{h} \frac{1}{m} \sum_{i}^{m} y_{i} h(x_{i})$$
(5)

$$\sigma_{i} = \begin{cases} +1 & \text{with prob .5} \\ -1 & \text{with prob .5} \end{cases}$$
(6)

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This gives us Rademacher correlation—what's the best that a random classifier could do?

$$\hat{\mathscr{R}}_{S}(H) \equiv \mathbb{E}_{\sigma} \left[ \max_{h \in H} \frac{1}{m} \sum_{i}^{m} \sigma_{i} h(x_{i}) \right]$$
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(7)

Notation:  $\mathbb{E}_{p}[f] \equiv \sum_{x} p(x) f(x)$ 

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$$\hat{\mathscr{R}}_{\mathcal{S}}(H) \equiv \mathbb{E}_{\sigma} \left[ \max_{h \in H} \frac{1}{m} \sum_{i}^{m} \sigma_{i} h(\mathbf{x}_{i}) \right]$$
(7)

Note: Empirical Rademacher complexity is with respect to a sample.

$$|H| = 1 \qquad \qquad |H| = 2^m$$

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$$|H| = 2^{m}$$

$$h(x_{i})\mathbb{E}_{\sigma}\left[\frac{1}{m}\sum_{i}^{m}\sigma_{i}\right] = 0$$

$$\frac{m}{m} = 1$$

$$|H| = 1$$

$$|H| = 2^{m}$$

$$h(x_{i})\mathbb{E}_{\sigma}\left[\frac{1}{m}\sum_{i}^{m}\sigma_{i}\right] = 0$$

$$\frac{|H| = 2^{m}}{\frac{m}{m}} = 1$$

- Rademacher correlation is larger for more complicated hypothesis space.
- What if you're right for stupid reasons?

We can generalize Rademacher complexity to consider all sets of a particular size.

$$\mathscr{R}_m(H) = \mathbb{E}_{S \sim D^m} \left[ \hat{\mathscr{R}}_S(H) \right] \tag{8}$$

#### Theorem

**Convergence Bounds** Let *F* be a family of functions mapping from *Z* to [0, 1], and let sample  $S = (z_1, ..., z_m)$  were  $z_i \sim D$  for some distribution *D* over *Z*. Define  $\mathbb{E}[f] \equiv \mathbb{E}_{z \sim D}[f(z)]$  and  $\hat{\mathbb{E}}_S[f] \equiv \frac{1}{m} \sum_{i=1}^m f(z_i)$ . With probability greater than  $1 - \delta$  for all  $f \in F$ :

$$\mathbb{E}[f] \le \hat{\mathbb{E}}_{s}[f] + 2\mathcal{R}_{m}(F) + \mathcal{O}\left(\sqrt{\frac{\ln\frac{1}{\delta}}{m}}\right)$$
(8)

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(8)

f is a surrogate for the accuracy of a hypothesis (mathematically convenient)

If we have a function:

$$|f(x_1,\ldots,x_i,\ldots,x_m)-f(x_1,\ldots,x_i',\ldots,x_m)| \le c_i$$
(9)

then:

$$\Pr[f(x_1,\ldots,x_m) \ge \mathbb{E}\left[f(X_1,\ldots,X_m)\right] + \epsilon] \le \exp\left\{\frac{-2\epsilon^2}{\sum_i^m c_i^2}\right\}$$
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Proofs online and in Mohri (requires Martingale, constructing  $V_k = \mathbb{E} [V | x_1 \dots x_k] - \mathbb{E} [V | x_1 \dots x_{k-1}]).$ 

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(10)

Proofs online and in Mohri (requires Martingale, constructing  $V_k = \mathbb{E} [V | x_1 \dots x_k] - \mathbb{E} [V | x_1 \dots x_{k-1}]).$ What function do we care about for Rademacher complexity? Let's define

$$\Phi(S) = \sup_{f} \left( \mathbb{E}[f] - \hat{\mathbb{E}}_{S}[f] \right) = \sup_{f} \left( \mathbb{E}[f] - \frac{1}{m} \sum_{i} f(z_{i}) \right)$$
(11)

#### Lemma

## Moving to Expectation With probability at least $1 - \delta$ , $\Phi(S) \leq \mathbb{E}_{s}[\Phi(S)] + \sqrt{\frac{\ln \frac{1}{\delta}}{2m}}$

Since  $f(z_1) \in [0, 1]$ , changing any  $z_i$  to  $z'_i$  in the training set will change  $\frac{1}{m} \sum_i f(z_i)$  by at most  $\frac{1}{m}$ , so we can apply McDiarmid's inequality with  $\epsilon = \sqrt{\frac{\ln \frac{1}{\delta}}{2m}}$  and  $c_i = \frac{1}{m}$ .

Define a ghost sample  $S' = (z'_1, ..., z'_m) \sim D$ . How much can two samples from the same distribution vary?

#### Lemma

#### **Two Different Samples**

$$\mathbb{E}_{S}[\Phi(S)] = \mathbb{E}_{S}\left[\sup_{f} (\mathbb{E}[f] - \hat{\mathbb{E}}_{S}[f])\right]$$
(12)  
(13)

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(12)

$$= \mathbb{E}_{\mathcal{S}}\left[\sup_{f \in \mathcal{F}} \left(\mathbb{E}_{\mathcal{S}'}\left[\hat{\mathbb{E}}_{\mathcal{S}'}\left[f\right]\right] - \hat{\mathbb{E}}_{\mathcal{S}}\left[f\right]\right)\right]$$
(13)

(14)

The expectation is equal to the expectation of the empirical expectation of all sets S'

#### Step 2: Comparing two different empirical expectations

Define a ghost sample  $S' = (z'_1, ..., z'_m) \sim D$ . How much can two samples from the same distribution vary?

#### Lemma

#### **Two Different Samples**

$$\mathbb{E}_{\mathcal{S}}[\Phi(\mathcal{S})] = \mathbb{E}_{\mathcal{S}}\left[\sup_{f} (\mathbb{E}[f] - \hat{\mathbb{E}}_{\mathcal{S}}[f])\right]$$
(12)

$$= \mathbb{E}_{S} \left[ \sup_{f \in F} (\mathbb{E}_{S'} \left[ \hat{\mathbb{E}}_{S'}[f] \right] - \hat{\mathbb{E}}_{S}[f]) \right]$$
(13)

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(15)

# S and S' are distinct random variables, so we can move inside the expectation

Define a ghost sample  $S' = (z'_1, ..., z'_m) \sim D$ . How much can two samples from the same distribution vary?

Lemma

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(13)

$$\leq \mathbb{E}_{\mathcal{S},\mathcal{S}'}\left[\sup_{f} \left(\hat{\mathbb{E}}_{\mathcal{S}'}[f] - \hat{\mathbb{E}}_{\mathcal{S}}[f]\right)\right]$$
(14)

The expectation of a max over some function is at least the max of that expectation over that function

From *S*, *S'* we'll create *T*, *T'* by swapping elements between *S* and *S'* with probability .5. This is still idependent, identically distributed (iid) from *D*. They have the same distribution:

$$\hat{\mathbb{E}}_{S'}[f] - \hat{\mathbb{E}}_{S}[f] \sim \hat{\mathbb{E}}_{T'}[f] - \hat{\mathbb{E}}_{T}[f]$$
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(15)

Let's introduce  $\sigma_i$ :

$$\hat{\mathbb{E}}_{T'}[f] - \hat{\mathbb{E}}_{T}[f] = \frac{1}{m} \begin{cases} f(z_i) - f(z'_i) \text{ with prob .5} \\ f(z'_i) - f(z_i) \text{ with prob .5} \end{cases}$$
(16)  
$$= \frac{1}{m} \sum_{i} \sigma_i (f(z'_i) - f(z_i))$$
(17)

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$$= \frac{1}{m} \sum_{i} \sigma_i (f(z'_i) - f(z_i))$$
(17)

Thus:  

$$\mathbb{E}_{\mathcal{S},\mathcal{S}'}\left[\sup_{f\in\mathcal{F}}\left(\hat{\mathbb{E}}_{\mathcal{S}'}\left[f\right]-\hat{\mathbb{E}}_{\mathcal{S}}\left[f\right]\right)\right] = \mathbb{E}_{\mathcal{S},\mathcal{S}',\sigma}\left[\sup_{f\in\mathcal{F}}\left(\sum_{i}\sigma_{i}(f(z'_{i})-f(z_{i}))\right)\right].$$

## Before, we had $\mathbb{E}_{S,S',\sigma} \left[ \sup_{f \in F} \sum_i \sigma_i (f(z'_i) - f(z_i)) \right]$

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$$\leq \mathbb{E}_{S,S',\sigma} \left[ \sup_{f \in F} \sum_{i} \sigma_{i} f(z'_{i}) + \sup_{f \in F} \sum_{i} (-\sigma_{i}) f(z_{i}) \right]$$
(18)
(19)

Taking the sup jointly must be less than or equal the individual sup.

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$$\leq \mathbb{E}_{S,S',\sigma} \left[ \sup_{f \in F} \sum_{i} \sigma_{i} f(z'_{i}) \right] + \mathbb{E}_{S,S',\sigma} \left[ \sup_{f \in F} \sum_{i} (-\sigma_{i}) f(z_{i}) \right]$$
(19)  
(20)

Linearity

Before, we had 
$$\mathbb{E}_{S,S',\sigma}\left[\sup_{f\in F}\sum_i \sigma_i(f(z'_i) - f(z_i))\right]$$

$$\leq \mathbb{E}_{S,S',\sigma} \left[ \sup_{f \in F} \sum_{i} \sigma_{i} f(z'_{i}) + \sup_{f \in F} \sum_{i} (-\sigma_{i}) f(z_{i}) \right]$$
(18)  
$$\leq \mathbb{E}_{S,S',\sigma} \left[ \sup_{f \in F} \sum_{i} \sigma_{i} f(z'_{i}) \right] + \mathbb{E}_{S,S',\sigma} \left[ \sup_{f \in F} \sum_{i} (-\sigma_{i}) f(z_{i}) \right]$$
(19)  
$$= \mathscr{R}_{m}(F) + \mathscr{R}_{m}(F)$$
(20)

Definition

$$\Phi(S) \leq \mathbb{E}_{S}[\Phi(S)] + \sqrt{\frac{\ln \frac{1}{\delta}}{2m}}$$

Step 1

$$\sup_{f} \left( \mathbb{E}\left[f\right] - \hat{\mathbb{E}}_{\mathcal{S}}\left[h\right] \right) \leq \mathbb{E}_{\mathcal{S}}\left[\Phi(\mathcal{S})\right] + \sqrt{\frac{\ln\frac{1}{\delta}}{2m}}$$

Definition of  $\Phi$ 

$$\mathbb{E}[f] - \hat{\mathbb{E}}_{\mathcal{S}}[h] \leq \mathbb{E}_{\mathcal{S}}[\Phi(\mathcal{S})] + \sqrt{\frac{\ln \frac{1}{\delta}}{2m}}$$

Drop the sup, still true

$$\mathbb{E}\left[f\right] - \hat{\mathbb{E}}_{\mathcal{S}}\left[h\right] \le \mathbb{E}_{\mathcal{S},\mathcal{S}'}\left[\sup_{f}\left(\hat{\mathbb{E}}_{\mathcal{S}'}\left[f\right] - \hat{\mathbb{E}}_{\mathcal{S}}\left[f\right]\right)\right] + \sqrt{\frac{\ln\frac{1}{\delta}}{2m}}$$
  
Step 2

$$\mathbb{E}[f] - \hat{\mathbb{E}}_{\mathcal{S}}[h] \le \mathbb{E}_{\mathcal{S},\mathcal{S}',\sigma} \left[ \sup_{f \in \mathcal{F}} \left( \sum_{i} \sigma_{i}(f(z_{i}') - f(z_{i})) \right) \right] + \sqrt{\frac{\ln \frac{1}{\delta}}{2m}}$$
(21)

Step 3

$$\mathbb{E}[f] - \hat{\mathbb{E}}_{\mathcal{S}}[h] \leq 2\mathscr{R}_m(F) + \sqrt{\frac{\ln \frac{1}{\delta}}{2m}}$$

Step 4

$$\mathbb{E}[f] - \hat{\mathbb{E}}_{\mathcal{S}}[h] \leq \frac{2\mathscr{R}_m(F)}{2m} + \sqrt{\frac{\ln\frac{1}{\delta}}{2m}}$$
(21)

Recall that  $\hat{\mathscr{R}}_{\mathcal{S}}(F) \equiv \mathbb{E}_{\sigma}\left[\sup_{f} \frac{1}{m} \sum_{i} \sigma_{i} f(z_{i})\right]$ , so we apply McDiarmid's inequality again (because  $f \in [0, 1]$ ):

$$\hat{\mathscr{R}}_{S}(F) \leq \mathscr{R}_{m}(F) + \sqrt{\frac{\ln \frac{1}{\delta}}{2m}}$$
(22)

**Putting the Pieces Together** 

With probability  $\geq 1 - \delta$ :

$$\mathbb{E}[f] - \hat{\mathbb{E}}_{\mathcal{S}}[h] \leq \frac{2\mathscr{R}_m(F)}{2m} + \sqrt{\frac{\ln\frac{1}{\delta}}{2m}}$$
(21)

Recall that  $\hat{\mathscr{R}}_{S}(F) \equiv \mathbb{E}_{\sigma}\left[\sup_{f} \frac{1}{m} \sum_{i} \sigma_{i} f(z_{i})\right]$ , so we apply McDiarmid's inequality again (because  $f \in [0, 1]$ ):

$$\hat{\mathscr{R}}_{S}(F) \leq \mathscr{R}_{m}(F) + \sqrt{\frac{\ln \frac{1}{\delta}}{2m}}$$
(22)

Putting the two together:

$$\mathbb{E}[f] \le \hat{\mathbb{E}}_{s}[f] + 2\mathscr{R}_{m}(F) + \mathscr{O}\left(\sqrt{\frac{\ln\frac{1}{\delta}}{m}}\right)$$
(23)

Define:

$$Z \equiv X \times \{-1, +1\} \tag{24}$$

$$f_h(x,y) \equiv \mathbb{1} \left[ h(x) \neq y \right] \tag{25}$$

$$F_H \equiv \{f_h : h \in H\}$$
(26)

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(25)

$$F_H \equiv \{f_h : h \in H\}$$
(26)

We can use this to create expressions for generalization and empirical error:

$$R(h) = \mathbb{E}_{(x,y)\sim D} \left[ \mathbb{1} \left[ h(x) \neq y \right] \right] = \mathbb{E} \left[ f_h \right]$$
(27)

$$\hat{R}(h) = \frac{1}{m} \sum_{i} \mathbb{1} \left[ h(x_i) \neq y \right] = \hat{\mathbb{E}}_{\mathcal{S}} \left[ f_h \right]$$
(28)

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(27)

$$\widehat{R}(h) = \frac{1}{m} \sum_{i} \mathbb{1} \left[ h(x_i) \neq y \right] = \widehat{\mathbb{E}}_{\mathcal{S}} \left[ f_h \right]$$
(28)

We can plug this into our theorem!

#### **Generalization bounds**

We started with expectations

$$\mathbb{E}[f] \le \hat{\mathbb{E}}_{\mathcal{S}}[f] + 2\hat{\mathscr{R}}_{\mathcal{S}}(F) + \mathcal{O}\left(\sqrt{\frac{\ln\frac{1}{\delta}}{m}}\right)$$
(29)

• We also had our definition of the generalization and empirical error:

$$R(h) = \mathbb{E}_{(x,y)\sim D}[\mathbb{1}[h(x)\neq y]] = \mathbb{E}[f_h] \quad \hat{R}(h) = \frac{1}{m}\sum_i \mathbb{1}[h(x_i)\neq y] = \hat{\mathbb{E}}_{\mathcal{S}}[f_h]$$

Combined with the previous result:

$$\hat{\mathscr{R}}_{\mathcal{S}}(F_{\mathcal{H}}) = \frac{1}{2}\hat{\mathscr{R}}_{\mathcal{S}}(\mathcal{H})$$
(30)

• All together:

$$R(h) \le \hat{R}(h) + \mathscr{R}_m(H) + \mathscr{O}\left(\sqrt{\frac{\log \frac{1}{\delta}}{m}}\right)$$
(31)

- Interaction of data, complexity, and accuracy
- Still very theoretical
- Next up: How to evaluate generalizability of specific hypothesis classes

Rademacher complexity provides nice guarantees

$$R(h) \leq \hat{R}(h) + \mathscr{R}_m(H) + \mathscr{O}\left(\sqrt{\frac{\log \frac{1}{\delta}}{2m}}\right)$$

- But in practice hard to compute for real hypothesis classes
- Is there a relationship with simpler combinatorial measures?

(32)

## Define the **growth function** $\Pi_H : \mathbb{N} \to \mathbb{N}$ for a hypothesis set *H* as:

$$\forall m \in \mathbb{N}, \Pi_H(m) \equiv \max_{\{x_1, \dots, x_m\} \in X} \left| \{ (h(x_1), \dots, h(x_m) : h \in H\} \right|$$
(33)

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(33)

i.e., the number of ways *m* points can be classified using *H*.

## If G is a function taking values in $\{-1, +1\}$ , then

$$\mathscr{R}_m(G) \le \sqrt{\frac{2\ln \Pi_G(m)}{m}} \tag{34}$$

Uses Masart's lemma





$$VC(H) \equiv \max\left\{m : \Pi_H(m) = 2^m\right\}$$
(35)





# $VC(H) \equiv \max\left\{m \colon \Pi_H(m) = 2^m\right\}$ (35)

The size of the largest set that can be fully shattered by *H*.

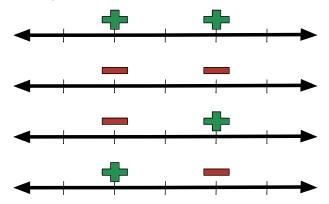
- Need upper and lower bounds
- Lower bound: example
- Upper bound: Prove that no set of *d* + 1 points can be shattered by *H* (harder)

• What about two points?

#### Intervals

What is the VC dimension of [a, b] intervals on the real line.

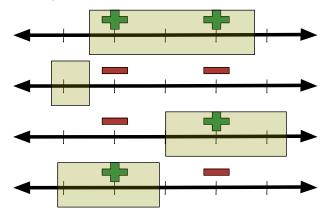
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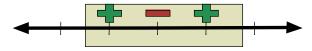
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- What about three points?
- **No set** of three points can be shattered
- Thus, VC dimension of intervals is 2

 Consider hypothesis that classifies points on a line as either being above or below a sine wave

$$\{t \to \sin(\omega x) : \omega \in \mathbb{R}\}$$
(36)

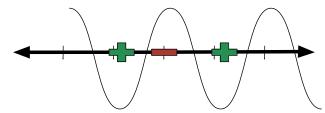
Can you shatter three points?

#### **Sine Functions**

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#### **Sine Functions**

 Consider hypothesis that classifies points on a line as either being above or below a sine wave

$$\{t \to \sin(\omega x) : \omega \in \mathbb{R}\}$$
(36)

• Can you shatter four points?

#### **Sine Functions**

 Consider hypothesis that classifies points on a line as either being above or below a sine wave

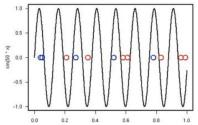
$$\{t \to \sin(\omega x) : \omega \in \mathbb{R}\}$$
(36)

• How many points can you shatter?

 Consider hypothesis that classifies points on a line as either being above or below a sine wave

$$\{t \to \sin(\omega x) : \omega \in \mathbb{R}\}$$
(36)

• Thus, VC dim of sine on line is  $\infty$ 



VC dimension obviously encodes the complexity of a hypothesis class, but we want to connect that to Rademacher complexity and the growth function so we can prove generalization bounds. VC dimension obviously encodes the complexity of a hypothesis class, but we want to connect that to Rademacher complexity and the growth function so we can prove generalization bounds.

### Theorem

**Sauer's Lemma** Let *H* be a hypothesis set with VC dimension *d*. Then  $\forall m \in \mathbb{N}$ 

$$\Pi_{H}(m) \leq \sum_{i=0}^{d} \binom{m}{i} \equiv \Phi_{d}(m)$$
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This is good because the sum when multiplied out becomes  $\binom{m}{i} = \frac{m \cdot (m-1) \dots}{i!} = \mathcal{O}(m^d)$ . When we plug this into the learning error limits:  $\log(\Pi_H(2m)) = \log(\mathcal{O}(m^d)) = \mathcal{O}(d \log m)$ .

# Prelim:

This comes from Pascal's Triangle This convention is consistent with Pascal's Triangle

### Prelim:

We'll proceed by induction. Our two base cases are:

- If m = 0,  $\Pi_H(m) = 1$ . You have no data, so there's only one (degenerate) labeling
- If d = 0,  $\Pi_H(m) = 1$ . If you can't even shatter a single point, then it's a fixed function

Assume that it holds for all m', d' for which m' + d' < m + d. We are given H, |S| = m,  $S = \langle x_1, \dots, x_m \rangle$ , and d is the VC dimension of H.

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 Build two new hypothesis spaces

  $\mathcal{H}$   $\mathcal{H}_1$   $\mathcal{H}_2$ 
 $\mathbf{x}_1, \dots, \mathbf{x}_m$   $\mathbf{x}_1, \dots, \mathbf{x}_{m-1}$  

 h1
  $\mathbf{0}$   $\mathbf{x}_1, \dots, \mathbf{x}_{m-1}$ 
 $\mathbf{h}_1$   $\mathbf{0}$   $\mathbf{x}_1, \dots, \mathbf{x}_{m-1}$ 
 $\mathbf{h}_1$   $\mathbf{0}$   $\mathbf{h}_1$   $\mathbf{h}_1$  <th

Encodes where the extended set has differences on the first *m* points.

$$|\Pi_{H}(S)| = |H_{1}| + |H_{2}|$$

$$\leq \sum_{n=1}^{d} \binom{m-1}{n} + \sum_{n=1}^{d-1} \binom{m-1}{n}$$
(38)

$$\leq \sum_{i=0} \left( \begin{array}{c} i \end{array} \right) + \sum_{i=0} \left( \begin{array}{c} i \end{array} \right)$$
(39)

(40)

$$|\Pi_{H}(S)| = |H_{1}| + |H_{2}|$$

$$\leq \sum_{i=0}^{d} {\binom{m-1}{i}} + \sum_{i=0}^{d-1} {\binom{m-1}{i}}$$
(39)
(40)

We can rewrite this as  $\sum_{i=0}^{d} {m-1 \choose i-1}$  because  ${x \choose -1} = 0$ .

$$|\Pi_{H}(S)| = |H_{1}| + |H_{2}|$$

$$< \sum_{n=1}^{d} {m-1 \choose n} + \sum_{n=1}^{d-1} {m-1 \choose n-1}$$
(38)
(39)

$$\leq \sum_{i=0}^{d} \left( \begin{array}{c} i \end{array} \right) + \sum_{i=0}^{d} \left( \begin{array}{c} i \end{array} \right)$$
(39)
$$= \sum_{i=0}^{d} \left[ \binom{m-1}{i} + \binom{m-1}{i-1} \right]$$
(40)

(41)

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(42)

## Pascal's Triangle

$$|\Pi_{H}(S)| = |H_{1}| + |H_{2}|$$
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$$\leq \sum_{i=0}^{d} {\binom{m-1}{i}} + \sum_{i=0}^{d-1} {\binom{m-1}{i}}$$
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$$= \sum_{i=0}^{d} \left[ {\binom{m-1}{i}} + {\binom{m-1}{i-1}} \right]$$
(40)  
$$= \sum_{i=0}^{d} {\binom{m}{i}}$$
(41)  
$$= \Phi_{d}(m)$$
(42)

Is this combinatorial expression really  $\mathcal{O}(m^d)$ ?

$$\begin{split} \sum_{i=0}^{d} \binom{m}{i} &\leq \sum_{i=0}^{d} \binom{m}{i} \left(\frac{m}{d}\right)^{d-i} \\ &\leq \sum_{i=0}^{m} \binom{m}{i} \left(\frac{m}{d}\right)^{d-i} \\ &= \left(\frac{m}{d}\right)^{d} \sum_{i=0}^{m} \binom{m}{i} \left(\frac{d}{m}\right)^{i} \\ &= \left(\frac{m}{d}\right)^{d} \left(1 + \frac{d}{m}\right)^{m} \leq \left(\frac{m}{d}\right)^{d} e^{d}. \end{split}$$

Combining our previous generalization results with Sauer's lemma, we have that for a hypothesis class *H* with VC dimension *d*, for any  $\delta > 0$  with probability at least  $1 - \delta$ , for any  $h \in H$ ,

$$R(h) \le \hat{R}(h) + \sqrt{\frac{2d\log\frac{em}{d}}{m}} + \sqrt{\frac{\log\frac{1}{\delta}}{2m}}$$
(43)

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- In class: more VC dimension examples